



Generalized Tensor Decompositions for Non-Normal Data

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ustration by Chris Brigma



A Tensor is an Multi-Way Array



 d^{th} -order Tensor

d > 3



Vector

d = 1



Tensors Come From Many Applications

- Chemometrics: Emission x Excitation x Samples (Fluorescence Spectroscopy)
- Neuroscience: Neuron x Time x Trial (Calcium Imaging)
- Criminology: Day x Hour x Location x Crime (Chicago Crime Reports)
- Medicine: Channel x Wavelength x Time (EEG measurements)
- Sports: Player x Statistic x Season
- Cyber-Traffic: IP x IP x Port x Time
- Social Network: Person x Person x Time x Interaction-Type



Tensor Decomposition: A Mathematical & Statistical Tool for Analysis of Tensor Data









Break Tensor into Understandable Parts...



Key: The parts have structure!





Given *d* vectors:

$$\mathbf{a}_k \in \mathbb{R}^{n_k}$$
 for $k = 1, \dots, d$

The **outer product** is

$$\mathbf{\mathcal{P}} = \mathbf{a}_1 \circ \mathbf{a}_2 \cdots \circ \mathbf{a}_d \in \mathbb{R}^{n_1 \times n_2 \times \cdots \times n_d}$$



$$\mathbf{\mathcal{P}}(i_1, i_2, i_3) = \mathbf{a}_1(i_1) \, \mathbf{a}_2(i_2) \, \mathbf{a}_3(i_3)$$

CANDECOMP/PARAFAC (CP) Tensor Factorization Uncovers the Rank-1 Parts





Hitchcock, 1927; Carroll and Chang, 1970; Harshman, 1970

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Frank Lauren Hitchcock

MIT Professor

(1875 - 1957)

CP first invented in 1927

THE EXPRESSION OF A TENSOR OR A POLYADIC AS A SUM OF PRODUCTS

By FRANK L. HITCHCOCK

1. Addition and Multiplication.

Tensors are added by adding corresponding components. The product of a covariant tensor $A_{i_1 \dots i_n}$ of order p into a covariant tensor $B_{i_{\alpha+1}} \dots i_{\alpha+\alpha}$ of order q is defined by writing

> $A_{i_1 \dots i_p} B_{i_p+1} \dots i_{p+q} = C_{i_1 \dots i_{p+q}}$ (1)

where the product $C_{i_1 \cdots i_{p+q}}$ is a covariant tensor of order p+q. When no confusion results indices may be omitted giving AB = C

 (1_{s})

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equivalent to the n^{p+q} equations (1). Boldface type is convenient for indicating that the letters do not denote merely numbers or scalars. Products of contravariant and of mixed tensors may be similarly defined.

A partial statement of the problem to be considered is as follows: to find under what conditions a given tensor can be expressed as a sum of products of assigned form. A more general statement of the problem will be given below.

2. Polyadic form of a tensor.

Any covariant tensor $A_{i_1..i_p}$ can be expressed as the sum of a finite number of tensors each of which is the product of p covariant vectors.

$$A_{i_{1}\cdots i_{p}} = \sum_{j=1}^{j=h} a_{lj, i_{1}} a_{2j, i_{2}} \cdots a_{pj, i_{p}}$$
(2)

where a_{1j, i}, etc., are a set of hp covariant vectors. When the indiccs $i_1 \cdot \cdot i_n$ can be omitted this may be written i = h

$$\mathbf{A} = \sum_{j=1}^{\Sigma} \mathbf{a}_{1j} \mathbf{a}_{2j} \cdot \cdot \mathbf{a}_{pj}. \tag{2}$$

The right member is now identical in appearance with a Gibbs

F. L. Hitchcock, *The Expression of a Tensor or* a Polyadic as a Sum of Products, Journal of Mathematics and Physics, 1927

2. Polyadic form of a tensor.

Any covariant tensor $A_{i_1 \dots i_p}$ can be expressed as the sum of a finite number of tensors each of which is the product of p covariant vectors.

$$A_{i_{1}} \dots i_{p} = \sum_{j=1}^{j=h} a_{ij, i_{1}} a_{2j, i_{2}} \cdots a_{pj, i_{p}}$$
(2)

where $a_{ij,i}$, etc., are a set of hp covariant vectors. When the indices $i_1 \cdot i_p$ can be omitted this may be written

$$\mathbf{A} = \sum_{j=1}^{j=h} \mathbf{a}_{1j} \mathbf{a}_{2j} \cdot \cdot \mathbf{a}_{pj}. \tag{2a}$$





CP Independently Reinvented (twice) in 1970

CANDECOMP: <u>Can</u>onical <u>Decomp</u>osition

PSYCHOMETRIKA-VOL. 35, NO. 3 BEPTEMBER, 1970

ANALYSIS OF INDIVIDUAL DIFFERENCES IN MULTIDIMEN-SIONAL SCALING VIA AN N-WAY GENERALIZATION OF "ECKART-YOUNG" DECOMPOSITION

J. DOUGLAS CARROLL AND JIH-JIE CHANG

BELL TELEPHONE LABORATORIES MURRAY HILL, NEW JERSEY

An individual differences model for multidimensional scaling is outlined in which individuals are assumed differentially to weight the several dimensions of a common "psychological space". A corresponding method of analyzing similarities data is proposed, involving a generalisation of "Eckart-Young analysis" to decomposition of three-way (or higher-way) tables. In the present case this decomposition is optical to a derived threeway table of scalar products between stimuli for individuals. This analysis yields a stimulus by dimensions coordinate matrix and a subjects by dimensions matrix of weights. This method is illustrated with data on auditory stimuli and on perception of nations.

There has been an interest for some time in the question of dealing with individual differences among subjects in making similarity judgments on which a multidimensional scaling of stimuli is to be based. Kruskal [1968] and McGee [1968] have both incorporated different ways of dealing with individual differences into their scaling procedures. Tucker and Messick [1963] proposed an approach, which they called "Points of view analysis," which is probably the most widely used method for dealing with such individual differences. In this method, intercorrelations are first computed between subjects (based on their similarity judgments) and the resulting correlation matrix is factor analyzed to produce a subject space. One then looks for clusters of subjects in this subject space, and if such clusters are found, proceeds in one way or another to define "idealized" subjects corresponding to clusters. (The "idealized subject" for a given cluster may be defined, for example, by finding the pattern of similarity judgments corresponding to a hypothetical subject at the cluster centroid, by choosing the actual subject closest to that centroid, or, most simply, by averaging the similarity judgments for subjects in the given cluster.) The similarities for these "idealized subjects" are then, individually and independently, subjected to multidimensional scaling.

This approach has been criticized by a number of people, most recently by Ross [1966] (see Cliff, 1968, for a reply to Ross's criticism and a further discussion of the "idealized individuals" interpretation of "Points of view

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Richard A. Harshman Univ. Ontario (1943-2008)

CP: CANDECOMP/PARAFAC

In 2000, Henk Kiers proposed this *compromise* name



2010: Pierre Comon, Lieven DeLathauwer, and others reverse-engineered CP, revising some of Hitchcock's terminology

PARAFAC: <u>Para</u>llel <u>Fac</u>tors

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NOTE: This manuscript was originally published in 1970 and is reproduced here to make it more accessible to interested scholars. The original reference is Harshman, R. A. (1970). Foundations of the PARFAC procedure: Models and conditions for an "explanatory" multimodal factor analysis. UCLA Working Papers in Phonetics, 16, 1-84. (University Microfilms, Ann Arbor, Michigan, No. 10,085).

FOUNDATIONS OF THE PARAFAC PROCEDURE: MODELS AND CONDITIONS

FOR AN "EXPLANATORY" MULTIMODAL FACTOR ANALYSIS



Many thanks to the following persons for helping me learn about Jih-Jie Chang: Fan Chung, Ron Graham, Shen Lin (husband), May Chang (niece), Lili Bruer (daughter).

Hitchcock, 1927; Carroll and Chang, 1970; Harshman, 1970

Standard CP: Sum of Squares Error (SSE)

 \wedge

 \square

x	$\approx \mathcal{M} = \left[\begin{array}{c} + \\ + \\ \end{array} \right] + \cdots + $	_
Ctondard CD	$\begin{array}{l l} \min \ F(\mathbf{X}, \mathbf{M}) \equiv \sum_{i \in \Omega} (x_i - m_i)^2 \\ \text{s.t. } \operatorname{rank}(\mathbf{M}) \leq r \end{array}$	

 \wedge

Shorthand for element of data tensor: $x_i \equiv x(i_1, i_2, \dots, i_d)$

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Element of model low-rank tensor: $m_i \equiv \sum_{j=1}^r \prod_{k=1}^d \mathbf{A}_k(i_k, j)$

(defined in terms of factor matrices)

 Ω = set of all n^d elements in tensor

Generalized CP (GCP)





by
$$\min F(\mathbf{X}, \mathbf{M}) \equiv \sum_{i \in \Omega} f(x_i, m_i)$$
s.t. rank $(\mathbf{M}) \leq r$

Why?

 SSE: maximum likelihood estimate (MLE) for Gaussian distribution

$$x_i = m_i + \epsilon, \ \epsilon \sim \mathcal{N}(0, \sigma)$$

 $x_i \sim \mathcal{N}(m_i, \sigma)$

- Different MLEs for different distributions
 - Poisson (counts)
 - Bernoulli (binary)

Probability Distribution ⇒ Maximum Likelihood Estimator





Gaussian MLE (Standard CP)

PDF for Normal Distribution
$$p(x | \mu, \sigma) = \frac{e^{-(x-\mu)^2/2\sigma^2}}{\sqrt{2\pi\sigma^2}}$$
 and
Link Function
 $m = \mu$
 σ constantNegative log-likelihood: $-\log p(x | \mu, \sigma) = \frac{(x-u)^2}{2\sigma^2} + \frac{1}{2}\log(2\pi\sigma^2)$ Eliminate natural parameter
via link function: $f(x,m) = \frac{(x-m)^2}{2\sigma^2} + \frac{1}{2}\log(2\pi\sigma^2)$ Eliminate constants: $f(x,m) = (x-m)^2$



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Hong, Kolda, Duersch, SIAM Review, 2019

PDF 1



Odds (m)

1/4

1

4

10

Bernoulli MLE with Odds Link (Binary Data)

Bernoulli random variable

$$x \in \{0,1\}$$

 $\rho = \text{probability of a 1}$
 $p(x \mid \rho) = \rho^x (1 - \rho)^{(1-x)}, \quad x \in \{0,1\}$
 $\ell(\rho) = \rho / (1 - \rho)$
 $\ell^{-1}(m) = m / (1 + m)$

Probability (ρ)

20%

50%

80%

90%

PMF for Bernoulli Distribution
$$p(x \mid \rho) = \rho^x (1 - \rho)^{(1 - x)}$$
 and
$$x \in \{0, 1\}$$

Link Function $m = \frac{\rho}{(1-\rho)}$

Negative log-likelihood:

$$-\log p(x \mid \rho) = \log \frac{1}{1 - \rho} - x \log \frac{\rho}{1 - \rho}$$

Eliminate natural parameter via link function:

 $f(x,m) = \log(1+m) - x\log m \quad \text{for} \quad m > 0$





Bernoulli MLE with Odds Link (Binary Data)



PMF for Bernoulli Distribution
$$p(x \mid \rho) = \rho^x (1 - \rho)^{(1 - x)}$$
 and
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Eliminate natural parameter via link function:

$$f(x,m) = \log(1+m) - x\log m \quad \text{for} \quad m > 0$$



Bernoulli MLE with Logit Link (Binary Data)

Bernoulli random variable

$$x \in \{0,1\}$$

 $\rho = \text{probability of a 1}$
 $p(x \mid \rho) = \rho^x (1 - \rho)^{(1-x)}, \quad x \in \{0,1\}$
 $\log(\rho) = \log(\rho / (1 - \rho))$
 $\log(\rho + \rho) = \log(\rho - \rho)$

Log-Odds(m)	Probability ($ ho$)
-1.39	20%
0	50%
1.39	80%
2.30	90%

PMF for Bernoulli Distribution $p(x \mid \rho) = \rho^x (1 - \rho)^{(1 - x)} \text{ and } x \in \{0, 1\}$

Link Function $m = \log \frac{\rho}{(1-\rho)}$

Negative log-likelihood:

$$-\log p(x \mid \rho) = \log \frac{1}{1 - \rho} - x \log \frac{\rho}{1 - \rho}$$

Eliminate natural parameter via link function:

 $f(x,m) = \log(1+e^m) - xm \text{ for } m \in \mathbb{R}$





Bernoulli MLE with Logit Link (Binary Data)



PMF for Bernoulli Distribution
$$p(x \mid \rho) = \rho^x (1 - \rho)^{(1 - x)}$$
 and
$$x \in \{0, 1\}$$

Link Function
$$m = \log \frac{\rho}{(1-\rho)}$$

Negative log-likelihood:

$$-\log p(x \mid \rho) = \log \frac{1}{1 - \rho} - x \log \frac{\rho}{1 - \rho}$$

Eliminate natural parameter via link function:

$$f(x,m) = \log(1+e^m) - xm \quad \text{for} \quad m \in \mathbb{R}$$



Sampling of Loss Functions





Example Tensor from Neuroscience

Source: Williams, et al. Unsupervised Discovery of Demixed, Low-dimensional Neural Dynamics across Multiple Timescales through Tensor Components Analysis. Neuron, 2018. <u>https://doi.org/10.1016/j.neuron.2018.05.015</u>

Activity of Single Neuron Measured Over Time Produces Vector Data



Thanks to Schnitzer Group @ Stanford Mark Schnitzer, Fori Wang, Tony Kim

111 time bins



Williams et al., Neuron, 2018

Multiple Neurons Measured Over Time Produces Matrix

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Thanks to Schnitzer Group @ Stanford Mark Schnitzer, Fori Wang, Tony Kim

Microscope by Inscopix



mouse in "maze"

neural activity



282 neurons \times 111 time bins



Williams et al., Neuron, 2018



Multiple Trials Produces 3-way Tensor





282 neurons \times 111 time bins \times 300 trials

Williams et al., Neuron, 2018

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Example Neuron Activity





Neuron Factor Vector Visualized as Bar Chart





Time Factor Vector Visualized as Line



Trial Factor Vector Visualized as Color-Coded Scatter Plot





Visualization of CP Tensor Decomposition Shows the Factors (Vectors)







"Standard" CP Decomposition of Mouse Data, aka Gaussian ($f(x,m) = (x-m)^2$)









CP Tensor Decomposition "Sees" Reward



CP Tensor Decomposition "Sees" Turn Direction





CP Tensor Decomposition Can be Tough to Interpret due to Negative Entries







GCP Decomposition with Beta Divergence ($\beta = 0.5, f(x, m) = \sqrt{m} + x/\sqrt{m}$)



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Regression Using GCP Factors on Trial Mode

Trial Factor Matrix is 300×8

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)	50	100	150	200	250	30

Look at predicting turn and reward. Split into two groups of 150 trials. Train regression model with 1st group. Test with 2nd group. Repeat 100 times.

$\min_{\boldsymbol{\beta}} \ \mathbf{A}_3^{\mathrm{train}} \boldsymbol{\beta} - \mathbf{y}^{\mathrm{train}}\ $
$\mathbf{\hat{y}}^{\text{test}} = \begin{bmatrix} \mathbf{A}_3^{\text{test}} \boldsymbol{\beta} \ge 0.5 \end{bmatrix}$





Optimization Formulation for GCP Tensor Decomposition





$$\begin{array}{l|l} \textbf{GO} & \min \ F(\mathbf{X}, \mathbf{M}) \equiv \sum_{i \in \Omega} f(x_i, m_i) \\ \text{s.t. } \operatorname{rank}(\mathbf{M}) \leq r & & i = \text{multi-index} \\ \Omega = \text{all indices} \end{array}$$

 Standard CP [Hitchcock, 1927; Carrol & Chang, 1970; Harshman, 1970]

 $f(x,m) = (x-m)^2$

 Poisson CP (Identity Link) [Welling & Webber, 2001; Chi & Kolda, 2009]

 $f(x,m) = m - x \log m$

 Logistic CP, etc. [Hong, Kolda, Duersch, 2018]

$$f(x,m) = \log(m+1) - x\log(m)$$

 \mathfrak{X} = \mathfrak{M} \approx d-way data *d*-way low-rank rank-one rank-one rank-one tensor of model tensor of size component component component size n^d n^d and rank ri = 1i = 2i = r $\mathfrak{X} \approx \mathfrak{M}$ where $\mathfrak{M} = \sum \mathbf{A}_1(:,j) \circ \mathbf{A}_2(:,j) \circ \cdots \circ \mathbf{A}_d(:,j)$ $\operatorname{rank}(\mathbf{M}) \leq r \ll n^d$ Low-rank: $\mathbf{A}_k \in \mathbb{R}^{n_k \times r}$ for $k \in \{1, \ldots, d\}$ Factor matrices: WLOG, $n = n_1 = \cdots = n_d$

Gradient-based Optimization for Fitting the GCP Model



s.t. $\operatorname{rank}(\mathbf{M}) \leq r$

C D

(J

<u>Define</u>: Elementwise partial gradient tensor, same size as data tensor = n^d

$$\mathbf{y} \qquad \qquad \mathbf{y}_i = \frac{\partial f}{\partial m}(x_i, m_i)$$

<u>Define</u>: Khatri-Rao product in all modes but one of size $n^{d-1} \times r$

$$\mathbf{Z}_k = \mathbf{A}_d \odot \cdots \odot \mathbf{A}_{k+1} \odot \mathbf{A}_{k-1} \odot \cdots \odot \mathbf{A}_1$$

Gradients computed via a sequence of matricizedtensor times Khatri-Rao product (MTTKRPs):

$$\mathbf{G}_{k} \equiv \frac{\partial F}{\partial \mathbf{A}_{k}} = \mathbf{Y}_{(k)} \mathbf{Z}_{k} \text{ for } k = 1, \dots, d \qquad \mathbf{MTTKRP}$$

$$\underset{k \text{ factor matrix of size } n \times r}{\text{ factor matrix of size } n \times r}$$

$$\mathbf{MTTKRPs \text{ can be computed efficiently...}}$$

$$\mathbf{MTTKRPs \text{ can be computed efficiently...}$$

$$\mathbf{MTTKRPs \text{ can be computed efficiently...}$$

$$\mathbf{MTTKRPs \text{ can be computed efficiently...}$$

$$\mathbf{MTTKRPs \text{ c$$

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Stochastic Gradient Descent (SGD) for GCP



Cost: O(rs) flops



$\mathbf{G}_k = \mathbf{Y}_{(k)} \mathbf{Z}_k$ Cost: $O(rn^d)$ flops Standard gradient $\min F(x)$ $y_i = \frac{\partial f}{\partial m}(x_i, m_i)$ y $x^{(t+1)} = x^{(t)} - \alpha q^{(t)}$ Stochastic gradient $ilde{\mathbf{G}}_k = ilde{\mathbf{Y}}_{(k)} \mathbf{Z}_k$ Choose stochastic *sparse* Y-tensor $\mathbb{E}[\tilde{\mathcal{Y}}] = \mathcal{Y}$ Ϋ́ such that $\operatorname{nnz}(\tilde{\mathbf{\mathcal{Y}}}) \leq s \ll n^d$

By linearity of expectation: $\mathbb{E}[\tilde{\mathbf{G}}_k] = \mathbf{G}_k$

Gradient Descent (GD) $\alpha =$ learning rate

Stochastic Gradient Descent (SGD) $x^{(t+1)} = x^{(t)} - \alpha \tilde{q}^{(t)}$ $\mathbb{E}[\tilde{q}^{(t)}] = q^{(t)} \equiv \nabla F(x^{(t)})$

Adam (Kingma & Ba, 2015) Adaptive momentum SGD

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Uniform Sampling



<u>Goal</u>: Random *sparse* tensor of size n^d that equals the "Y-tensor" in expectation

Sample
$$s \ll n^d$$
 random tensor
entries (with replacement)
 $\tilde{s}_i = \#$ times i sampled
 $\tilde{y}_i = \tilde{s}_i \cdot \frac{n^d}{s} \cdot y_i$
 $y_i = \frac{\partial f}{\partial m}(x_i, m_i)$
Claim: $\mathbb{E}[\tilde{y}] = y$

Proof: $\mathbb{E}[\tilde{s}_i] = \frac{s}{n^d}$ $\mathbb{E}[\tilde{y}_i] = \mathbb{E}[\tilde{s}_i] \cdot \frac{n^d}{s} \cdot y_i = y_i$



Choosing *s*, the number of sampled elements...

- Choose s = O(n)
- Gradient = O(rs) = O(rn) versus $O(rn^d)$ Downside...
- If data tensor is sparse, few entries corresponding to nonzeros will be chosen

Stratified 0/1 Sampling



ies

<u>Goal</u>: Random *sparse* tensor of size n^d that equals the "Y-tensor" in expectation



Semi-Stratified 0/1 Sampling





<u>Goal</u>: Random *sparse* tensor of size n^d that equals the "Y-tensor" in expectation



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GCP with Stochastic Optimization



- Nonconvex problem
 - No guarantees of finding minimizer
- Using Adam (Kingma & Ba, 2015)
 - Default parameters
 - Some tweaks for checking convergence
- Past work on recommender systems uses SGD but ignores zeros
 - Gemulla, Nijkamp, Hass, Sismanis, KDD'11
 - Zhuang, Chin, Juan, and Lin, RecSys'13
- Past work on streaming uses SGD but data appears one slice at a time
 - Mardani, Mateos, Giannakis, IEEE TSP 2015
 - Maehara, Hayashi, Kawarabayashi,



Example on Gamma-Distributed Data

 $200 \times 150 \times 100 \times 50$ Tensor with low-rank (r = 5) structure based on Gamma distribution ($k = 1, \theta$ from model). Gamma loss: $f(x, m) = \frac{x}{m} + \log m$. Running stochastic GCP with 25 random starts and varying numbers of samples.



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Stochastic vs. Non-Stochastic





Example on Bernoulli-Distributed Data



 $200 \times 150 \times 100 \times 50$ Tensor with low-rank (r = 5) structure based on Bernoulli distribution (odds from model). Sparse tensor, less than 0.35% dense (~500K nonzeros). Bernoulli loss: $f(x, m) = \log(m + 1) - x \log m$. Running stochastic GCP with 25 random starts, varying # of samples. 3.1 samples=125 Success at Recovering Dashed lines: Individual runs, Solid lines: Median, 3.05 samples=250 **Underlying Generative** samples) Epoch: Asterisk (success), Dot(fail). samples=500 samples=1000 Factors samples=2000 25 •••••nominal (true solution) 2.95 recoveries 05 estimated loss (100,000 2.9 solution 2.85 15 of true 2.8 10 2.75 number 5 2.7 250 500 1000 2000 125 2.65 gradient samples 20 10 30 40 50 60 70 80 time (sec)

Uniform Sampling is Worse than Stratified for Sparse Tensors





Chicago Crime Data

- 4-way count tensor
 - 6,186 Days
 - 24 Hours of the Day
 - 77 Community Areas
 - 32 Crime Types
- Non-zeros: 5,330,673
 - Storage: 0.21GB for sparse tensor
- Distribution of entries
 - 0: 98.54%
 - **1**: 1.33%
 - ≥ 2:0.12%
- Obtained from FROSTT (<u>http://frostt.io/tensors/chicago-crime/</u>)
- Data originally from Chicago Data Portal (<u>https://data.cityofchicago.org/Public-</u> <u>Safety/Crimes-2001-to-present/ijzp-q8t2</u>)



GCP-Count

Rank = 10

s = 6,319



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Application to Sparse Crime Binary Tensor (Semi-stratified results)





Component #1





0

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Component #3







Component #6



0

Aside: Estimating Higher-Order Moments via Symmetric Tensor Factorization



Joint work with Sam Sherman, Notre Dame

First-order moment (mean): $\frac{1}{p} \sum_{i=1}^{p} \mathbf{a}_{i}$ We can compute lowrank ($r \ll p$) Second-order moment: $\frac{1}{p} \sum_{i=1}^{p} \mathbf{a}_i \circ \mathbf{a}_i$ symmetric tensor estimated to higherorder moments... $\frac{1}{r}\sum_{i=1}^{r}\mathbf{b}_{i}\circ\mathbf{b}_{i}\circ\mathbf{b}_{i}$ Third-order moment: $\frac{1}{p} \sum_{i=1}^{p} \mathbf{a}_i \circ \mathbf{a}_i \circ \mathbf{a}_i$ Fourth-order moment: $\frac{1}{p} \sum_{i=1}^{p} \mathbf{a}_{i} \circ \mathbf{a}_{i} \circ \mathbf{a}_{i} \circ \mathbf{a}_{i} = \mathbf{a}_{i} \cdot \mathbf{a}_{i} = \mathbf$

Given a set of p observations: $\mathbf{a}_i \in \mathbb{R}^n, i = 1, 2, \dots, p$





What are good applications, if any?

Submit your work at simods.siam.org

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NOW, X=X=A W-1



References & Collaborators

My department is hiring statisticians! Talk to me to learn more.

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